

Stochastic State Space Models for Mortality

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*"Stochastic Period and Cohort Effect State-Space Mortality Models
Incorporating Demographic Factors via Probabilistic Robust Principle
Components"*

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Mortality Modelling Context

The modelling and management of systematic mortality risk are two of the main concerns of large life insurers and pension plans:

Modelling:

- ▶ What is the best way to *forecast future mortality rates* and to model the *uncertainty surrounding these forecasts*?
- ▶ How do we value risky future cashflows that depend on future mortality rates?

Management:

- ▶ How can this risk be actively managed and reduced as part of an overall strategy of efficient risk management?
- ▶ What hedging instruments are easier to price than others?

Mortality Modelling Context

Stylized Facts of Mortality Data: Enhancing mortality models requires an understanding of common features of mortality behaviour (see discussion in [Cairns, Blake and Dowd, 2008])

- ▶ Mortality rates have fallen dramatically at all ages.
- ▶ Rate of decrease in mortality has **varied over time and by age group**.
 - ▶ *For example, for English and Welsh males, the age 25 rate improved dramatically before 1960 and then levelled off; conversely at age 65 the opposite was true*
- ▶ Absolute decreases have **varied by age group**.
 - ▶ *For example, for English and Welsh males, the age 45 improvements have been much higher than the age 85 improvements.*
- ▶ Aggregate mortality rates have significant **volatility year on year**.

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Stochastic Mortality Models

The uncertainty in future death rates can be divided into two components:

- ▶ **Unsystematic mortality risk.** Even if the true mortality rate is known, the number of deaths, $D(t, x)$, will be random.
 - ▶ larger population \Rightarrow smaller unsystematic mortality risk (due to pooling of offsetting risks - diversification).
- ▶ **Systematic mortality risk.** This is the undiversifiable component of mortality risk that affects all individuals in the same way.
 - ▶ **Forecasts of mortality rates in future years are uncertain.**

Background on Stochastic Mortality Modelling

- ▶ Pricing of retirement income products depends crucially on the accuracy of the predicted death or survival probabilities.
- ▶ *It has been widely documented that **survival probability is consistently underestimated** especially in the last few decades ([IMF, 2012]).*
- ▶ To capture the stochastic nature of mortality trends, [Lee and Carter, 92] proposed a stochastic mortality model to forecast the trend of age-specific mortality rates.
- ▶ **Since the introduction of the Lee-Carter model, a range of stochastic mortality models have been proposed in the literature.**


Stochastic Mortality Models

Single age group models:

- ▶ Model the individual age group mortality evolution either: **force of mortality**¹ or **annual death counts**.
- ▶ Typically such models include:
 - ▶ *temporal smoothing splines*;
 - ▶ *demographic factors*;
 - ▶ can be *count processes* or *functional regressions* (or both);
 - ▶ *ARIMA* type structures.

Term structure of mortality (multiple age group) models:

- ▶ Typically model the **log mortality rate across the term structure of mortality**.
- ▶ Typically such models include:
 - ▶ *temporal smoothing splines*;
 - ▶ *period effects*; and
 - ▶ *cohort effects*.

¹force of mortality represents the instantaneous rate of mortality at a certain age measured on an annualized basis. It is identical in concept to failure rate, also called hazard function, in reliability theory. 

Stochastic Mortality Models: Regression Formulations

Generalized Linear Model Type: have been widely adopted in mortality modelling (eg. [Forfar, 1988], [Renshaw, 2003,2000,1991], [Currie,2016]).

Modelling target was either:

- ▶ the probability of death q_x , based on initial exposures; or
- ▶ the force of mortality μ_x , based on central exposures.

When targeting q_x it was common to use a **binomial observation distribution**

When targeting the force of mortality it was common practice to use a **Poisson observation distribution**

Stochastic Mortality Models: Time-series Regression

Stochastic Period Effect Models: influential stochastic factor model for mortality modelling given by [Lee and Carter, 1992]

- ▶ Dynamics of the log crude death rates, $y_{x,t} = \ln \hat{m}_{x,t}$, follow:

$$y_{x,t} = \alpha_x + \beta_x \kappa_t + \varepsilon_{x,t}, \quad \varepsilon_{x,t} \stackrel{iid}{\sim} N(0, \sigma_\varepsilon^2),$$

- ▶ $\alpha = \alpha_{x_1:x_p} := [\alpha_{x_1}, \dots, \alpha_{x_p}]$ represents the *age-profile of the log death rates*
 - ▶ $\beta = \beta_{x_1:x_p}$ measures the *sensitivity of of death rates for different age group* to a change of period effect κ_t .
- ▶ The **period effect**, κ_t , *for forecasting*, is typically set as

$$\kappa_t = \kappa_{t-1} + \theta + \omega_t, \quad \omega_t \stackrel{iid}{\sim} N(0, \sigma_\omega^2),$$

where $\varepsilon_{x,t}$ and ω_t are independent.

Stochastic Mortality Models: Time-series Regression

Extensions to the LC model:

Model	Dynamics
[Lee and Carter,92]	$\ln(\hat{m}_{x,t}) = \alpha_x + \beta_x \kappa_t + \varepsilon_{x,t}$
[Renshaw et al, 03]	$\ln(\hat{m}_{x,t}) = \alpha_x + \sum_{i=1}^k \beta_x^{(i)} \kappa_t^{(i)} + \varepsilon_{x,t}$
[Renshaw et al, 06]	$\ln(\hat{m}_{x,t}) = \alpha_x + \beta_x^{(1)} \kappa_t + \beta_x^{(2)} \zeta_{t-x} + \varepsilon_{x,t}$
[Currie,06]	$\ln(\hat{m}_{x,t}) = \alpha_x + \kappa_t + \zeta_{t-x} + \varepsilon_{x,t}$
[Cairns et al.,06]	$\text{logit}(q_{x,t}) = \kappa_t^{(1)} + \kappa_t^{(2)}(x - \bar{x})$
[Cairns et al., 09]	$\text{logit}(q_{x,t}) = \kappa_t^{(1)} + \kappa_t^{(2)}(x - \bar{x}) + \zeta_{t-x}$
[Plat, 09]	$\ln(\hat{m}_{x,t}) = \alpha_x + \kappa_t^{(1)} + \kappa_t^{(2)}(\bar{x} - x) + \kappa_t^{(3)}(\bar{x} - x)^+ + \zeta_{t-x} + \varepsilon_{x,t}$

Recently, [Fung et al, 16] and [Fung et al, 17] have proposed new extensions based on stochastic volatility structures in the latent processes as well as non-homoscedasticity in the mortality term structures, long-memory persistence and demographic factor model distributed lags.

Background on Stochastic Mortality Modelling

Frequentist Models with demographic and economic data.

- ▶ In [Hanewald,2011] and [Niu,2014] investigate links between *economic growth and morality trends* via regression model:
 - ▶ **Single Age Group Period effect Lee-Carter model + covariate (GDP).**
- ▶ [Hanewald,2011] included *cause-of-death* categorical variables.
- ▶ [Murray and Lopez, 1997] developed a multi factor linear regression model: the *logarithm of the rate of mortality per age group, sex and clustered cause of death* is regressed against the *socio-economic, educational, technological and cause-of-death* related regressors.
- ▶ [Hyndman and Yasmeen, 2012] and [Erbas, 2010] considered dimension reduction based feature extraction methods for regressors: *functional PCA covariates from mortality curves*

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State Space Based Stochastic Mortality Models

A key advantage of state space modelling is that the typical two-stage estimation and forecasting procedure under the SVD or Poisson regression maximum likelihood approaches can be combined in a single setting. This has the following advantages:

- ▶ more numerically and statistically robust than standard two stage regression modelling;
- ▶ can remove awkward identification specifications;
- ▶ is computationally more efficient; and
- ▶ can produce more accurate in-sample and out-of-sample forecasts;
- ▶ can be optimal from an efficiency and unbiased estimation perspective;
- ▶ easily adapted to Bayesian inference!

State Space Based Stochastic Mortality Models

State-Space Formulation: Observation Process

Let $y_x = \ln \tilde{m}_{x,t}$, in matrix notation we have (recall that $\gamma_t^x := \gamma_{t-x}$)

$$\begin{pmatrix} y_{x_1,t} \\ y_{x_2,t} \\ \vdots \\ y_{x_p,t} \end{pmatrix} = \begin{pmatrix} \alpha_{x_1} \\ \alpha_{x_2} \\ \vdots \\ \alpha_{x_p} \end{pmatrix} + \begin{pmatrix} \beta_{x_1} & \beta_{x_1}^\gamma & 0 & \cdots & 0 \\ \beta_{x_2} & 0 & \beta_{x_2}^\gamma & \cdots & 0 \\ \vdots & \vdots & \vdots & \ddots & \vdots \\ \beta_{x_p} & 0 & 0 & \cdots & \beta_{x_p}^\gamma \end{pmatrix} \begin{pmatrix} \kappa_t \\ \gamma_t^{x_1} \\ \gamma_t^{x_2} \\ \vdots \\ \gamma_t^{x_p} \end{pmatrix} + \begin{pmatrix} \varepsilon_{x_1,t} \\ \varepsilon_{x_2,t} \\ \vdots \\ \varepsilon_{x_p,t} \end{pmatrix}.$$

It is clear that, we have for $i \in \{1, \dots, p\}$:

$$y_{x_i,t} = \alpha_{x_i} + \beta_{x_i} \kappa_t + \beta_{x_i}^\gamma \gamma_t^{x_i} + \varepsilon_{x_i,t}$$

Here $(\kappa_t, \gamma_t^{x_1}, \dots, \gamma_t^{x_p})^\top$ is the $p + 1$ dimensional latent state vector.

State Space Based Stochastic Mortality Models

State-Space Formulation: State Process

An extended latent cohort dynamic for $\gamma_t^{x_1}$ is obtained by specifying the second row of the $p + 1$ by $p + 1$ matrix.

For example, one can consider generally the state equation as

$$\begin{pmatrix} \kappa_t \\ \gamma_t^{x_1} \\ \gamma_t^{x_2} \\ \vdots \\ \gamma_t^{x_{p-1}} \\ \gamma_t^{x_p} \end{pmatrix} = \begin{pmatrix} 1 & 0 & 0 & \cdots & 0 & 0 \\ 0 & \lambda_1 & \lambda_2 & \cdots & \lambda_{p-1} & \lambda_p \\ 0 & 1 & 0 & \cdots & 0 & 0 \\ 0 & 0 & 1 & \cdots & 0 & 0 \\ \vdots & \vdots & \vdots & \ddots & \vdots & \vdots \\ 0 & 0 & 0 & \cdots & 1 & 0 \end{pmatrix} \begin{pmatrix} \kappa_{t-1} \\ \gamma_{t-1}^{x_1} \\ \gamma_{t-1}^{x_2} \\ \vdots \\ \gamma_{t-1}^{x_{p-1}} \\ \gamma_{t-1}^{x_p} \end{pmatrix} + \begin{pmatrix} \theta \\ \eta \\ 0 \\ \vdots \\ 0 \\ 0 \end{pmatrix} + \begin{pmatrix} \omega_t^{\kappa} \\ \omega_t^{\gamma} \\ 0 \\ \vdots \\ 0 \\ 0 \end{pmatrix},$$

where

$$\gamma_t^{x_1} = \lambda_1 \gamma_{t-1}^{x_1} + \lambda_2 \gamma_{t-1}^{x_2} + \cdots + \lambda_{p-1} \gamma_{t-1}^{x_{p-1}} + \lambda_p \gamma_{t-1}^{x_p} + \eta + \omega_t^{\gamma}$$

which is an ARIMA(p,0,0) process since $\gamma_{t-1}^{x_i} = \gamma_{t-i}^{x_1}$, $i = 2, \dots, p$.

State Space Based Stochastic Mortality Models

Cohort effects: state-space formulation

We can express the matrix form succinctly as

$$\mathbf{y}_t = \boldsymbol{\alpha} + B\boldsymbol{\varphi}_t + \boldsymbol{\varepsilon}_t, \quad \boldsymbol{\varepsilon}_t \stackrel{iid}{\sim} N(0, \sigma_\varepsilon^2 \mathbf{1}_p),$$
$$\boldsymbol{\varphi}_t = \Lambda\boldsymbol{\varphi}_{t-1} + \boldsymbol{\Theta} + \boldsymbol{\omega}_t, \quad \boldsymbol{\omega}_t \stackrel{iid}{\sim} N(0, \Upsilon),$$

where

$$B = \begin{pmatrix} \beta_{x_1} & \beta_{x_1}^\gamma & 0 & \cdots & 0 \\ \beta_{x_2} & 0 & \beta_{x_2}^\gamma & \cdots & 0 \\ \vdots & \vdots & \vdots & \ddots & \vdots \\ \beta_{x_p} & 0 & 0 & \cdots & \beta_{x_p}^\gamma \end{pmatrix}, \quad \Lambda = \begin{pmatrix} 1 & 0 & 0 & \cdots & 0 & 0 \\ 0 & \lambda & 0 & \cdots & 0 & 0 \\ 0 & 1 & 0 & \cdots & 0 & 0 \\ 0 & 0 & 1 & \cdots & 0 & 0 \\ \vdots & \vdots & \vdots & \ddots & \vdots & \vdots \\ 0 & 0 & 0 & \cdots & 1 & 0 \end{pmatrix}, \quad \boldsymbol{\Theta} = \begin{pmatrix} \theta \\ \eta \\ 0 \\ \vdots \\ 0 \\ 0 \end{pmatrix},$$

and $\boldsymbol{\varphi}_t = (\kappa_t, \gamma_t^{x_1}, \dots, \gamma_t^{x_p})^\top$, $\mathbf{1}_p$ the p -dimensional identity matrix and Υ is a $p+1$ by $p+1$ diagonal matrix with diagonal $(\sigma_\kappa^2, \sigma_\gamma^2, 0, \dots, 0)$.

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State-Space Hybrid Factor Models for Mortality

- ▶ Extending stochastic mortality models with observable exogenous features/covariates from demographic data.
- ▶ This offers two advantages to standard Lee-Carter models:
 - ▶ *firstly they may improve predictive power of the models*
 - ▶ *secondly they may improve interpretation of the dynamic of the “term-structure” of age specific mortality rates.*

We address four new and important aspects in practice previously ignored:

1. **missing data** in time-series and panel (matrix) structured real demographic data;
2. **noisy observations and outliers** (in real data);
3. **parsimonious model** creation via dimension reduction; and
4. **optimal estimation** via **computational efficient** state-space filtering methods.

State-Space Hybrid Factor Models for Mortality

“Hybrid”: a mix of observable stochastic features and latent stochastic factors

- ▶ Two fundamental approaches to develop Hybrid Factor Models:
 1. **time varying factor** with **static loading coefficient**
(classical distributed lag regressions such as ARDL models);
 2. **static factor** with **time varying stochastic loading coefficients**.
(state space models e.g. dynamic Nelson-Siegel yield curves).
- ▶ Approach 2 is more appropriate for data which is **high dimensional** in nature, **time series / panel structured** but represented by relatively **“short time series”** lengths.
 - ▶ *This type of data is particularly prevalent in demographic studies!*
- ▶ The main concept here is that the **feature extraction** is performed over the entire available time series of observable demographic data.

State-Space Hybrid Factor Models for Mortality

- ▶ There are numerous ways to achieve this in a state-space model:
 - ▶ the *factor may influence all age groups equally* by entering the factor into the **state equation**; or
 - ▶ the *factor may influence each age specific mortality rate differently* by adding it in the **observation equation**.
- ▶ Denote generically \mathbf{F}_t as the $p \times k$ factors matrix where p may represent number of age groups and k may represent number of age specific factors.
- ▶ We then specify an additional latent pk dimensional vector ϱ_t for the factor loading for year t .
 - ▶ ϱ_t is a dynamic regression parameter for factors matrix \mathbf{F}_t which specifies the impact of $x_i \in \{x_1, \dots, x_p\}$ age group and $m \in \{1, \dots, k\}$ component corresponding to $[\mathbf{F}_t]_{i,m}$ by $\varrho_t^{i,m}$ element.

State-Space Hybrid Factor Models for Mortality

Consider the State-Space Hybrid Stochastic Period-Cohort-Demographic Model

The general notation of the model is as follows

$$\begin{aligned} \mathbf{y}_t &= \boldsymbol{\alpha} + \tilde{\mathbf{B}}_t \tilde{\boldsymbol{\varphi}}_t + \boldsymbol{\varepsilon}_t, \quad \boldsymbol{\varepsilon}_t \stackrel{iid}{\sim} \mathcal{N}(0, \sigma_\varepsilon^2 \mathbb{I}_p), \\ \tilde{\boldsymbol{\varphi}}_t &= \tilde{\mathbf{\Lambda}} \tilde{\boldsymbol{\varphi}}_{t-1} + \tilde{\boldsymbol{\Theta}} + \tilde{\boldsymbol{\omega}}_t, \quad \tilde{\boldsymbol{\omega}}_t \stackrel{iid}{\sim} \mathcal{N}(0, \tilde{\Upsilon}) \end{aligned}$$

where $\tilde{\boldsymbol{\varphi}}_t = (\boldsymbol{\varphi}_t, \boldsymbol{\varrho}_t)$ is a $(p + pk + 1) \times 1$ latent process vector and

$$\tilde{\boldsymbol{\Theta}} = \begin{pmatrix} \boldsymbol{\Theta}_{(p+1) \times 1} \\ \boldsymbol{\Psi}_{pk \times 1} \end{pmatrix}_{(p+pk+1) \times 1}$$

is a vector of drift parameters.

We assume independence of error terms in latent variables to give a covariance matrix for the state error $\tilde{\boldsymbol{\omega}}_t$:

$$\tilde{\Upsilon} = \left(\begin{array}{c|c} \Upsilon_{(p+1) \times (p+1)} & 0 \\ \hline 0 & \sigma_\varrho^2 \mathbb{I}_{pk} \end{array} \right)_{(p+pk+1) \times (p+pk+1)}$$

State-Space Hybrid Factor Models for Mortality

Define the following two objects: $\tilde{\mathbf{F}}_t = \bigoplus_{j=1}^k [\mathbf{F}_t]_{j,\cdot}$, and $\tilde{\mathbf{f}}_t = \text{vec}(\mathbf{F}_t^T)$ giving:

$$\tilde{\mathbf{F}}_t = \begin{pmatrix} [\mathbf{F}_t]_{1,\cdot} & 0 & 0 & \cdots & 0 \\ 0 & [\mathbf{F}_t]_{2,\cdot} & 0 & \cdots & 0 \\ \vdots & & \ddots & & \vdots \\ 0 & \cdots & & & [\mathbf{F}_t]_{p,\cdot} \end{pmatrix}_{p \times pk} \quad \text{and} \quad \tilde{\mathbf{f}}_t = \begin{pmatrix} [\mathbf{F}_t]_{1,1} \\ [\mathbf{F}_t]_{1,2} \\ \vdots \\ [\mathbf{F}_t]_{p,k} \end{pmatrix}_{pk \times 1}$$

where $[\mathbf{F}_t]_{j,\cdot}$ and $[\mathbf{F}_t]_{j,m}$ represent the j th row and the element corresponding to j th row and m th column of \mathbf{F}_t , respectively.

Consider three cases of model:

Case 1: *Factors in Observation Equation Only;*

Case 2: *Factors in Period Effect State Equation Only;*

Case 3: *Factors in Cohort Effect State Equation Only.*

State-Space Hybrid Factor Models for Mortality

We can now define the different sub-models:

$$\tilde{\mathbf{B}}_{t \ p \times (p+\rho k+1)} = \begin{cases} \left(\begin{array}{c|c} \mathbf{B}_{\rho \times (\rho+1)} & \tilde{\mathbf{F}}_t \end{array} \right) & \text{for Case 1,} \\ \left(\begin{array}{c|c} \mathbf{B}_{\rho \times (\rho+1)} & \mathbf{0}_{\rho \times \rho k} \end{array} \right) & \text{otherwise,} \end{cases}$$

$$\tilde{\mathbf{\Lambda}}_{(\rho+\rho k+1) \times (\rho+\rho k+1)} = \begin{cases} \left(\begin{array}{c|c} \mathbf{\Lambda}_{(\rho+1) \times (\rho+1)} & \mathbf{0}_{(\rho+1) \times \rho k} \\ \mathbf{0}_{\rho k \times (\rho+1)} & \mathbf{\Omega}_{\rho k \times \rho k} \end{array} \right) & \text{for Case 1,} \\ \left(\begin{array}{c|c} \mathbf{\Lambda}_{(\rho+1) \times (\rho+1)} & \tilde{\mathbf{f}}_t^T \\ \mathbf{0}_{\rho k \times (\rho+1)} & \mathbf{\Omega}_{\rho k \times \rho k} \end{array} \right) & \text{for Case 2,} \\ \left(\begin{array}{c|c} \mathbf{\Lambda}_{(\rho+1) \times (\rho+1)} & \mathbf{0}_{1 \times \rho k} \\ \mathbf{0}_{\rho k \times (\rho+1)} & \tilde{\mathbf{F}}_t \\ \mathbf{\Omega}_{\rho k \times \rho k} & \end{array} \right) & \text{for Case 3.} \end{cases}$$

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Mortality Data and Demographic Data Description

Country	Life Expectancy (E0)	No. Births	Death Rate (mx)	No. Deaths
Austria	1947 - 2014	1871 - 2014	1947 - 2014	1947 - 2014
Belarus	1959 - 2014	1959 - 2014	1959 - 2014	1959 - 2014
Belgium	1841 - 2015	1840 - 2015	1841 - 2015	1841 - 2015
Czech Republic	1950 - 2010	1947 - 2014	1950 - 2014	1950 - 2014
Denmark	1835 - 2014	1835 - 2014	1835 - 2014	1835 - 2014
Estonia	1959 - 2013	1959 - 2013	1959 - 2013	1959 - 2013
Finland	1878 - 2012	1865 - 2012	1878 - 2012	1878 - 2012
France	1816 - 2014	1806 - 2014	1816 - 2014	1816 - 2014
East Germany	1956 - 2013	1946 - 2013	1956 - 2013	1956 - 2013
West Germany	1956 - 2013	1946 - 2013	1956 - 2013	1956 - 2013
Greece	1981 - 2013	1981 - 2013	1981 - 2013	1981 - 2013
Estonia	1959 - 2013	1959 - 2013	1959 - 2013	1959 - 2013
Hungary	1950 - 2014	1950 - 2014	1950 - 2014	1950 - 2014
Iceland	1838 - 2013	1838 - 2013	1838 - 2013	1838 - 2013
Ireland	1950 - 2014	1950 - 2014	1950 - 2014	1950 - 2014

Table: Demographic data available per country (HM Database).

Mortality Data and Demographic Data Description

Country	Life Expectancy (E0)	No. Births	Death Rate (mx)	No. Deaths
Italy	1872 - 2012	1862 - 2012	1872 - 2012	1872 - 2012
Latvia	1959 - 2013	1959 - 2013	1959 - 2013	1959 - 2013
Lithuania	1959 - 2013	1959 - 2013	1959 - 2013	1959 - 2013
Luxembourg	1960 - 2014	1950 - 2014	1960 - 2014	1960 - 2014
Netherlands	1850 - 2012	1850 - 2012	1850 - 2012	1850 - 2012
Norway	1846 - 2014	1846 - 2014	1846 - 2014	1846 - 2014
Poland	1958 - 2014	1958 - 2014	1958 - 2014	1958 - 2014
Portugal	1940 - 2012	1886 - 2012	1940 - 2012	1940 - 2012
Russia	1959 - 2014	1959 - 2014	1959 - 2014	1959 - 2014
Slovakia	1950 - 2014	1950 - 2014	1950 - 2014	1950 - 2014
Slovenia	1983 - 2014	1983 - 2014	1983 - 2014	1983 - 2014
Spain	1908 - 2014	1908 - 2014	1908 - 2014	1908 - 2014
Sweden	1751 - 2014	1747 - 2014	1751 - 2014	1751 - 2014
Switzerland	1876 - 2014	1871 - 2014	1876 - 2014	1876 - 2014
United Kingdom	1922 - 2013	1922 - 2013	1922 - 2013	1922 - 2013
Ukraine	1959 - 2013	1946 - 2013	1959 - 2013	1959 - 2013

Table: Demographic data available per country (HH Database).

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Results and Analysis

The model estimation was performed by Forward-Backward Kalman Filter within Rao-Blackwellised Adaptive Gibbs Sampler (MCMC).

The models we considered in our studies were of type:

1. [LCC:] Lee-Carter model with the stochastic cohort effect.
2. [DFM-PC:] demographic factor model version of Lee-Carter (Period-Cohort).

The factors are obtained by performing **Probabilistic Principle Component Analysis PPCA** jointly on the set of data for all countries listed, excluding:

*United Kingdom (response variable)
Greece and Slovakia (due to short time series).*

Results and Analysis

[DFM-PC:] demographic factor model version of LCC sub-models:

- ▶ [DFM-PC-B:] the mean of first principal component of Birth counts as a static parameter, age specific element of ϱ_t ;
- ▶ [DFM-PC-D-r/s:] the first principal component of Death counts (which is age and country specific) as an exogenous factor, one element of ϱ_t corresponds to a country specific subvector of the component, robust standardisation (s = non-robust standardisation);
- ▶ [DFM-PC-Mx-r/s:] the first principal component of Death Rates (which is age and country specific) as an exogenous factor, one element of ϱ_t corresponds to a country specific subvector of the component, robust standardisation (s = non-robust standardisation);

Results and Analysis

Out-of-Sample Forecast Age Specific Log-Death Rates: Performance Analysis

- ▶ We choose for the out-of-sample study the last 10 years of the available sample for British Female death rates.
- ▶ Model calibration period is 1922 – 2002
⇒ forecast performance analysis for 2003 – 2013

Model	MSE	DIC	MSEP _{MCMC}	MSEP _{Kalman}
LCC	0.0097	-3627	0.1778	0.1774
DFM-PC-B	0.0072	-6500	0.0057	0.0062
DFM-PC-D-r	0.0182	-6380	0.0177	0.0251
DFM-PC-D-s	0.0065	-5996	0.0185	0.0156
DFM-PC-M _x -r	0.0081	-8225	0.0111	0.0129
DFM-PC-M _x -s	0.0174	-3951	0.0692	0.0285

- ▶ **The results confirm that adding demographic features, as additional explanatory variables to the LCC model, improves both in-sample fit out-of-sample fit and therefore the predictability of log death rates.**

Results and Analysis

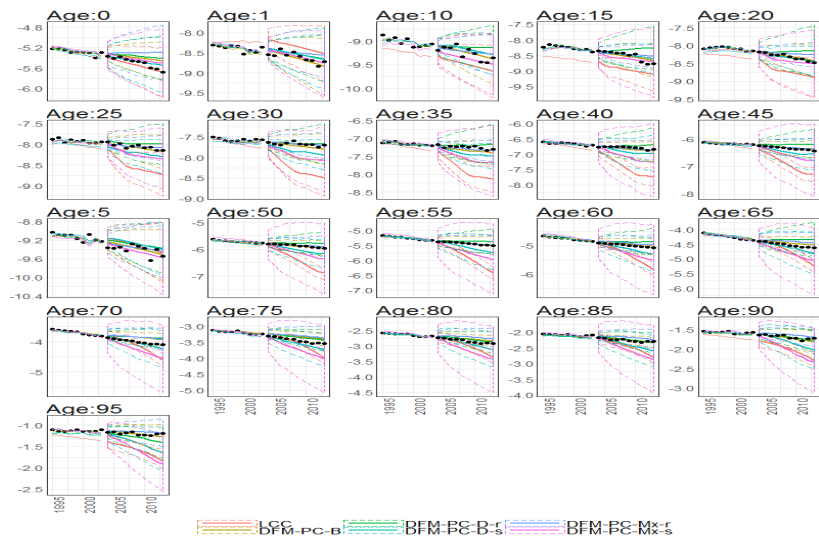


Figure: 10-year out-of-sample forecasted log death (y axis) rates by age with corresponding prediction intervals.

Conclusions

- ▶ We explored how to construct a state space formulation of the stochastic mortality models for Period and Cohort factors
- ▶ We explored how to extend to Hybrid Multi-Factor Stochastic State-Space Mortality models with Period-Cohort factors as well as demographic regressors.
- ▶ We briefly learnt about feature/covariate extraction methods to extract the demographic factors used in the extended HMF Stochastic State-Space Mortality models.
- ▶ We see that the standard Lee-Carter stochastic mortality models consistently under performs in-sample and out-of-sample in a range of estimation criteria, compared to the new proposed models.
- ▶ Lee-Carter Period-Cohort model consistently under estimates log-death rates
- ▶ Extended models proposed improve significantly the forecast performance of log-death rates.

Recent Papers on Stochastic Mortality Modelling

- 1 Fung M.C., Peters G.W., Shevchenko P.V.
A unified approach to mortality modelling using state-space framework: characterisation, identification, estimation and forecasting.
Annals of Actuarial Science. 2017 May:1-47.
Available at SSRN: <https://ssrn.com/abstract=2786559>
- 2 Fung M.C. and Peters G.W. and Shevchenko P.V.
A State-Space Estimation of the Lee-Carter Mortality Model and Implications for Annuity Pricing
MODSIM Modelling and Simulation Society. 2015, July.
Available at SSRN: <https://ssrn.com/abstract=2699624>
- 3 Toczydlowska D., Peters G.W., Fung M.C. and Shevchenko P.V.
Stochastic Period and Cohort Effect State-Space Mortality Models Incorporating Demographic Factors via Probabilistic Robust Principle Components Risks: Special Issue on "Aging Population Risks".
Available at SSRN: <https://ssrn.com/abstract=2977306>
- 4 Fung M.C., Peters G.W. and Shevchenko P.V.
Cohort Effects in Mortality Modelling: A Bayesian State-Space Approach (March 24, 2017).
Available at SSRN: <https://ssrn.com/abstract=2907868>